

ASYMPTOTIC RISK COMPARISON OF SOME ESTIMATORS FOR BIVARIATE NORMAL COVARIANCE MATRIX

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Abstract. For Selliah's [5] and Stein's [3] loss functions, Sharma and Krishnamoorthy [6] have considered orthogonal equivariant estimators Ψ . In this paper, we obtain asymptotic risk expressions for Ψ and make a numerical comparison with those for Haff's [2] and Sugiura-Fujimoto [7] estimators. It is observed that Ψ is uniformly better than Haff's estimator and better than Sugiura-Fujimoto estimator except in a small region of the parameter space.

1. Introduction.

Let S have a Wishart distribution with unknown matrix Σ and n degrees of freedom and the problem be one of estimating Σ under Selliah's [5] loss function $L_1(\Sigma, \hat{\Sigma}) = \text{tr}(\hat{\Sigma}\Sigma^{-1} - I)^2$ or Stein's [3] loss function $L_2(\Sigma, \hat{\Sigma}) = \text{tr}(\hat{\Sigma}\Sigma^{-1}) - \log|\hat{\Sigma}\Sigma^{-1}| - p$. The usual estimator of Σ is S/n . Through invariance considerations Selliah [5] and Stein [3] obtained the best lower triangular equivariant estimators φ_1 and φ_2 respectively which are better than any multiple aS of S . From Kiefer's [4] theorem their estimators are also minimax. Haff [2] using an identity for a Wishart matrix found an unbiased estimator of the difference between the risks of aS and the estimators proposed by him and hence proved that his estimators are better. Sugiura and Fujimoto [7] suggested some empirical Bayes estimators and derived asymptotic risk expressions for the best lower triangular equivariant estimators φ_i , Haff's [2] estimators and their own estimators. From their results, it is clear that neither Haff's nor Sugiura-Fujimoto estimators are uniformly better than φ_i . Following a suggestion by Eaton [1], Sharma and Krishnamoorthy [6] have obtained orthogonal equivariant estimator Ψ_i dominating φ_i under the loss L_i when Σ is 2×2 .

In this paper, we derive asymptotic risk expressions for Ψ_i ($i=1, 2$). Numerical comparison of the asymptotic risks in Sections 3 and 4 shows that Ψ_i is uniformly better than the corresponding Haff's estimator; this is consistent with the conclusions from the Monte-Carlo study in Sharma and Krishnamoorthy [ibid].

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It is also better than Sugiura-Fujimoto estimator except in a small region of the parameter space.

2. Asymptotic Risk Expressions of $\Psi_1(S)$ and $\Psi_2(S)$.

In this section, we derive asymptotic risk expressions of $\Psi_1(S)$ and $\Psi_2(S)$ under the loss functions L_1 and L_2 respectively. The estimator $\Psi_1(S)$ of Σ , when Σ is 2×2 is given by

$$(2.1) \quad \Psi_1(S) = a_1 S + b_1 \frac{S^{1/2}}{\text{tr } S^{-1/2}}$$

where, if $S = \Gamma D_\lambda \Gamma'$ with Γ orthogonal and $D_\lambda = \text{diag}(\lambda_1, \lambda_2)$, $S^{1/2} = \Gamma D_\lambda^{1/2} \Gamma'$ with $D_\lambda^{1/2} = \text{diag}(\sqrt{\lambda_1}, \sqrt{\lambda_2})$, $S^{-1/2} = (S^{1/2})^{-1}$, and

$$a_1 = (n^2 + n + 2)/(n^3 + 5n^2 + 6n + 4), \quad b_1 = 2n/(n^3 + 5n^2 + 6n + 4).$$

To find the risk expression of $\Psi_1(S)$, we need the following lemma. For the proof, see Sharma and Krishnamoorthy [6, Sec. 3].

LEMMA 2.1. *Let $S \sim W_2(\Sigma, n)$, Then*

$$E \left[\frac{|S|^{1/2}}{\text{tr } S^{1/2}} \text{tr}(S^{1/2} \Sigma^{-1}) \right] = (n-1)$$

$$E \left[\frac{|S|^{1/2}}{\text{tr } S^{1/2}} \text{tr}(S \Sigma^{-1} S^{1/2} \Sigma^{-1}) \right] = (n+2)(n-1)$$

$$E \left[\frac{|S|}{(\text{tr } S^{1/2})^2} \text{tr}(S^{1/2} \Sigma^{-1} S^{1/2} \Sigma^{-1}) \right] = (n+1) \left(n E \frac{\text{tr } S}{(\text{tr } S^{1/2})^2} - 1 \right).$$

Denoting the risk under L_i by R_i and using Lemma 2.1, we have

$$(2.2) \quad \begin{aligned} R_1(\Sigma, \Psi_1(S)) - R_1(\Sigma, a_1 S) \\ = b_1^2 (n+1) \left[n E \frac{\text{tr } S}{(\text{tr } S^{1/2})^2} - 1 \right] + 2b_1 (n-1) [a_1 (n+2) - 1]. \end{aligned}$$

To find the asymptotic expansion of $E[(\text{tr } S)/(\text{tr } S^{1/2})^2]$, we need the following identity used in Sugiura and Fujimoto [ibid]:

For any analytic function $f(S)$,

$$(2.3) \quad E f(S/n) = f(\Sigma) + n^{-1} \text{tr}(\Sigma \partial) f(A) |_{A=\Sigma} + O(n^{-2})$$

where ∂ is a matrix of differential operators and its (i, j) element is given by $(1/2)(1 + \delta_{ij}) \left(\frac{\partial}{\partial \lambda_{ij}} \right)$ for $A = (\lambda_{ij})$ and $\delta_{ij} = 1$ or 0 according as $i = j$ or $i \neq j$. It is convenient to write

$$(2.4) \quad \frac{1}{n} \text{tr}(\Sigma \partial)^2 f(A)|_{A=\Sigma} = \frac{1}{n} \Sigma \sigma_{rs} \sigma_{tu} \partial_{st} \partial_{ur} f(A)|_{A=\Sigma},$$

so that the operator ∂ acts on $f(A)$ only.

Since, for S of order 2×2 ,

$$(\text{tr } S)/(\text{tr } S^{1/2})^2 = (1 + 2|S|^{1/2}/\text{tr } S)^{-1},$$

we have, from (2.2) and (2.3),

$$(2.5) \quad E \frac{\text{tr } S}{(\text{tr } S^{1/2})^2} = \left(1 + \frac{2|\Sigma|^{1/2}}{\text{tr } \Sigma}\right)^{-1} + \frac{1}{n} \Sigma \sigma_{rs} \sigma_{tu} \partial_{st} \partial_{ur} \left(1 + \frac{2|A|^{1/2}}{\text{tr } A}\right)^{-1} |_{A=\Sigma} + O(n^{-2}).$$

After evaluating the second term on the right hand side of (2.5), it can be seen that

$$(2.6) \quad E \frac{\text{tr } S}{(\text{tr } S^{1/2})^2} = \left(1 + \frac{2|\Sigma|^{1/2}}{\text{tr } \Sigma}\right)^{-1} + \frac{|\Sigma|^{1/2}}{n \text{tr } \Sigma} \left(1 + \frac{2|\Sigma|^{1/2}}{\text{tr } \Sigma}\right)^{-2} \times \left[4 - 4\left(1 + \frac{2|\Sigma|^{1/2}}{\text{tr } \Sigma}\right)^{-1} \frac{|\Sigma|^{1/2}}{\text{tr } \Sigma} + 8\left(1 + \frac{2|\Sigma|^{1/2}}{\text{tr } \Sigma}\right)^{-1} \times \frac{|\Sigma|^{1/2} \text{tr } \Sigma^2}{(\text{tr } \Sigma)^3} - 4 \frac{\text{tr } \Sigma^2}{(\text{tr } \Sigma)^2}\right] + O(n^{-2}).$$

Since Ψ_1 is orthogonal and scale equivariant take $\Sigma = \text{diag}(1, c)$, $c > 0$, without loss of generality. Let $c^* = (1 + 2\sqrt{c}/(1+c))^{-1}$. Then, (2.6) can be written as

$$(2.7) \quad E \frac{\text{tr } S}{(\text{tr } S^{1/2})^2} = c^* + \frac{1}{n} \left(\frac{\sqrt{c}}{1+c}\right) (c^*)^2 \left[4 - 4c^* \frac{\sqrt{c}}{1+c} \left(1 - 2 \frac{(1+c)^2}{(1+c)^2}\right) - 4 \frac{1+c^2}{(1+c)^2}\right] + O(n^{-2}).$$

Thus, using (2.7) and

$$R_1(\Sigma, a_1 S) = 2n(n+3)a_1^2 - 4na_1 + 2,$$

we get from (2.2)

$$(2.8) \quad R_1(\Sigma, \Psi_1(S)) = 2n(n+3)a_1^2 - 4na_1 + 2 + 2b_1(n-1)[a_1(n+2) - 1] + b_1^2(n+1)(n \times (\text{approx.}) - 1) + O(n^{-4}),$$

where (approx.) is the expression given in (2.7).

Next, let us find the asymptotic risk expression of $\Psi_2(S)$ under L_2 . For $\Psi_2(S)$ given as

$$a_2 S + b_2 \frac{S^{1/2}}{\text{tr } S^{-1/2}} \quad \text{with } a_2 = 1/(n+1) \quad \text{and } b_2 = 2/(n^2-1),$$

in Sharma and Krishnamoorthy [ibid], we have

$$(2.9) \quad R_2(\Sigma, \Psi_2(S)) = E \operatorname{tr} \left(a_2 S + \frac{b_2 S^{1/2}}{\operatorname{tr} S^{-1/2}} \right) \Sigma^{-1} - E \log \left| \left(a_2 S + \frac{b_2 S^{1/2}}{\operatorname{tr} S^{-1/2}} \right) \Sigma^{-1} \right| - 2 \\ - a_2 E \operatorname{tr} S \Sigma^{-1} + b_2 E \frac{1}{\operatorname{tr} S^{-1/2}} \operatorname{tr} S^{1/2} \Sigma^{-1} - 2 \log a_2 \\ - E \log |S \Sigma^{-1}| - E \log |I + (b_2/a_2) S^{-1/2} / \operatorname{tr} S^{-1/2}| - 2.$$

Using the relations

$$E \log |S \Sigma^{-1}| = \sum_{j=1}^n E \log (\chi_{n-j+1}^2)$$

and

$$(2.10) \quad E \log |I + (b_2/a_2) S^{-1/2} / \operatorname{tr} S^{-1/2}| \\ = (b_2/a_2) - [b_2^2/(2a_2^2)] E \operatorname{tr} S / (\operatorname{tr} S^{1/2})^2 + O(n^{-3})$$

and Lemma 2.1, (2.9) can be written as

$$(2.11) \quad R_2(\Sigma, \Psi_2(S)) = 2 \log(n+1) - \sum_{j=1}^n E \log (\chi_{n-j+1}^2) \\ - b_2/a_2 + [b_2^2/(2a_2^2)] E \operatorname{tr} S / (\operatorname{tr} S^{1/2})^2.$$

Using digamma function $\zeta(x) = \Gamma'(x)/\Gamma(x)$, we have

$$(2.12) \quad \sum_{j=1}^n E \log (\chi_{n-j+1}^2) = 2 \log 2 + \sum_{j=1}^n \zeta \left(\frac{n-j+1}{2} \right).$$

Thus, from (2.11), (2.12) and (2.7),

$$(2.13) \quad R_2(\Sigma, \Psi_2(S)) = 2 \log \left(\frac{n+1}{2} \right) - \zeta(n/2) - \zeta \left(\frac{n-1}{2} \right) - 2/(n-1) \\ + \frac{2}{(n-1)^2} \times (\text{approx.}) + O(n^{-4}),$$

where, as before, (approx.) is the expression given in (2.7).

3. Comparison of the Estimators under the Loss L_1 .

Haff [2] has considered estimators of $\Sigma(p \times p)$ of the form

$$(3.1) \quad \Sigma_{1H} = a \left(S + \frac{b}{\operatorname{tr} S^{-1}} I \right), \quad a = (n+p+1)^{-1}, \quad 0 < b \leq 2(p-1)/(n-p+3),$$

which dominate the best multiple aS of S . Sugiura and Fujimoto [ibid] have derived an asymptotic risk expression of Σ_{1H} under the loss L_1 :

$$(3.2) \quad R_1(\Sigma, \Sigma_{1H}) - R_1(\Sigma, aS) \\ = \frac{2b}{(n+p+1)^2} \left[-n(p+1) + \left\{ 2n-4p-4 - bn(p+1) + \frac{bn^2}{2} \right\} \frac{\operatorname{tr} \Sigma^{-2}}{(\operatorname{tr} \Sigma^{-1})^2} \right]$$

$$+(p+1)^2 - \frac{8 \operatorname{tr} \Sigma^{-3}}{(\operatorname{tr} \Sigma^{-1})^3} + 3(bn+4) \frac{(\operatorname{tr} \Sigma^{-2})^2}{(\operatorname{tr} \Sigma^{-1})^4} \Big] + 0(n^{-4}).$$

The asymptotic value of b which minimizes an upper bound of the expression (3.2) is $b_0 = (p-1)(n+p-3)/n^2$. Also, we know that $R_1(\Sigma, aS) = \frac{p(p+1)}{n+p+1}$. Thus, for $p=2$ and $b=b_0$, when $\Sigma = \operatorname{diag}(1, c)$,

$$R_1(\Sigma, \Sigma_{1H}) = \frac{6}{n+3} + \frac{2(n-1)}{n^2(n+3)^2} \left[-3n + \frac{(5n-1)(n-6)}{2n} \cdot \frac{(1+c^2)}{(1+c)^2} \right. \\ \left. + 9 - \frac{8(1+c^3)}{(1+c)^3} + \frac{3(5n-1)}{n} \cdot \frac{(1+c^2)^2}{(1+c)^4} \right] + 0(n^{-4}).$$

Notice that Σ_{1H} is orthogonal and scale equivariant like Ψ_i , so that Σ can be taken to be $\operatorname{diag}(1, c)$ without loss of generality.

Sugiura and Fujimoto [ibid], on empirical Bayes arguments, suggested an estimator of the form

$$(3.4) \quad \Sigma_{1SF} = a \left(S + \frac{b \operatorname{tr} S^{-1} C}{\operatorname{tr}(S^{-1} C)^2} C \right),$$

where $a = (n+p+1)^{-1}$, $0 < b \leq 2(b-1)/n$ and C any positive definite matrix. It is asymptotically better than aS and if $b=0(n^{-1})$ and $C=I$,

$$(3.5) \quad R_1(\Sigma, \Sigma_{1SF}) - R_1(\Sigma, aS) \\ = \frac{2b}{(n+p+1)^2} \left[-2n + n \left(\frac{bn}{2} - p - 1 \right) E \frac{(\operatorname{tr} S^{-1})^2}{\operatorname{tr} S^{-2}} + 4n E \frac{\operatorname{tr} S^{-1} \operatorname{tr} S^{-3}}{(\operatorname{tr} S^{-2})^2} + 4p + 6 \right. \\ \left. + \{(p+1)^2 - 6 - bn(2p+3)/2\} \frac{(\operatorname{tr} \Sigma^{-1})^2}{\operatorname{tr} \Sigma^{-2}} - 16 \frac{\operatorname{tr} \Sigma^{-4}}{(\operatorname{tr} \Sigma^{-2})^2} \right. \\ \left. - 4(bn+2p+4) \frac{\operatorname{tr} \Sigma^{-1} \operatorname{tr} \Sigma^{-3}}{(\operatorname{tr} \Sigma^{-2})^2} + 32 \frac{\operatorname{tr} \Sigma^{-1} \operatorname{tr} \Sigma^{-5}}{(\operatorname{tr} \Sigma^{-2})^3} \right. \\ \left. + 8bn \frac{(\operatorname{tr} \Sigma^{-1})^2 \operatorname{tr} \Sigma^{-4}}{(\operatorname{tr} \Sigma^{-2})^3} - \frac{bn(\operatorname{tr} \Sigma^{-1})^4}{(\operatorname{tr} \Sigma^{-2})^2} \right] + 0(n^{-4}).$$

The asymptotic value of b , which minimizes an upper bound of the above expression, is $b_0 = (p-1)(n+p-3)/n^2$. Also, they established

$$(3.6) \quad E \frac{(\operatorname{tr} S^{-1})^2}{\operatorname{tr} S^{-2}} = (\operatorname{tr} \Sigma^{-1})^2 / \operatorname{tr} \Sigma^{-2} + \frac{1}{n} \left[\frac{8 \operatorname{tr} \Sigma^{-4} (\operatorname{tr} \Sigma^{-1})^2}{(\operatorname{tr} \Sigma^{-2})^3} - \frac{(\operatorname{tr} \Sigma^{-1})^4}{(\operatorname{tr} \Sigma^{-2})^2} \right. \\ \left. - \frac{8 \operatorname{tr} \Sigma^{-3} \operatorname{tr} \Sigma^{-1}}{(\operatorname{tr} \Sigma^{-2})^2} - \frac{(\operatorname{tr} \Sigma^{-1})^2}{\operatorname{tr} \Sigma^{-2}} + 2 \right] + 0(n^{-2}) = SF1(\text{say}).$$

and

$$(3.7) \quad E \left(\frac{\operatorname{tr} S^{-1} \operatorname{tr} S^{-3}}{(\operatorname{tr} S^{-2})^2} \right) = \frac{\operatorname{tr} \Sigma^{-1} \operatorname{tr} \Sigma^{-3}}{(\operatorname{tr} \Sigma^{-2})^2} + \frac{1}{n} \left[24 \frac{\operatorname{tr} \Sigma^{-1} \operatorname{tr} \Sigma^{-3} \operatorname{tr} \Sigma^{-4}}{(\operatorname{tr} \Sigma^{-2})^4} \right.$$

$$\begin{aligned}
& -\frac{2}{(\text{tr } \Sigma^{-2})^3} \{(\text{tr } \Sigma^{-1})^3 \text{tr } \Sigma^{-3} + 12 \text{tr } \Sigma^{-1} \text{tr } \Sigma^{-5} + 4(\text{tr } \Sigma^{-3})^2\} \\
& + \frac{1}{(\text{tr } \Sigma^{-2})^2} \{ \text{tr } \Sigma^{-1} \text{tr } \Sigma^{-3} + 6 \text{tr } \Sigma^{-4} \} + \frac{3(\text{tr } \Sigma^{-1})^2}{\text{tr } \Sigma^{-2}} \Big] + 0(n^{-2}) \\
& = SF2(\text{say}),
\end{aligned}$$

using (2.3).

As Σ_{1SF} is orthogonal and scale equivariant, we can take Σ to be a matrix of the form $\text{diag}(1, c_2, \dots, c_p)$. For $p=2$, $b=b_0$, $C=I$, and $\Sigma=\text{diag}(1, c)$, (3.2)-(3.7) imply

$$\begin{aligned}
(3.8) \quad R_1(\Sigma, \Sigma_{1SF}) &= R_1(\Sigma, aS) + \frac{2(n-1)}{n^2(n+3)^2} \\
& \times \left[-2n - \frac{5n+1}{2}(SF1) + 4n(SF2) + 14 - \frac{n-7}{2n} \cdot \frac{(1+c)^2}{1+c^2} \right. \\
& \quad + \{ -16(1+c^4)(1+c^2) - 4(9n-1)(1+c)(1+c^2)(1+c^3)/n \\
& \quad \quad + 32(1+c)(1+c^5) + 8(n-1)(1+c)^2(1+c^4)/n \\
& \quad \quad \left. - (n-1)(1+c)^4(1+c^2)/n \} / (1+c^2)^3 \right] + 0(n^{-4}).
\end{aligned}$$

We have done numerical calculation of the asymptotic risks of Ψ_1 , Σ_{1H} and Σ_{1SF} for various values of c and n . The Tables 1(a)-(e) are for $c \geq 1$; scale equivariance of the estimators implies that the values at c and c^{-1} are the same. The estimator Ψ_1 is seen to be uniformly better than Σ_{1H} and better than Σ_{1SF} except at $c=1, 2$ for $n \geq 20$.

Tables 1(a)-(e). Asymptotic comparison of the estimators Ψ_1 , Σ_{1H} and Σ_{1SF} . Columns (1), (2) and (3) give the values of $R_1(\Sigma, \Psi_1)$, $R_1(\Sigma, \Sigma_{1H}) - R_1(\Sigma, \Psi_1)$ and $R_1(\Sigma, \Sigma_{1SF}) - R_1(\Sigma, \Psi_1)$ respectively for $\Sigma = \text{diag}(1, c)$.

(a); $n=10$				(b); $n=20$			
c	(1)	(2)	(3)	c	(1)	(2)	(3)
1	.44414	.00203	-.00721	1	.25441	.00072	-.00316
2	.44438	.00258	-.00608	2	.25450	.00101	-.00151
3	.44473	.00336	-.00114	3	.25463	.00138	.00046
4	.44505	.00403	.00252	4	.25475	.00167	.00163
5	.44533	.00457	.00472	5	.25487	.00189	.00228
10	.44636	.00603	.00793	10	.25526	.00243	.00317
20	.44747	.00677	.00815	20	.25568	.00266	.00317
30	.44811	.00689	.00781	30	.25592	.00267	.00303
40	.44854	.00686	.00749	40	.25608	.00264	.00291
50	.44887	.00679	.00723	50	.25620	.00261	.00281
100	.44976	.00642	.00643	100	.25653	.00245	.00252
200	.45049	.00597	.00574	200	.25680	.00227	.00226

(c); $n=40$				(d); $n=80$			
c	(1)	(2)	(3)	c	(1)	(2)	(3)
1	.13754	.00023	-.00100	1	.07173	.00007	-.00028
2	.13757	.00033	-.00034	2	.07174	.00010	-.00008
3	.13761	.00046	.00028	3	.07175	.00013	.00010
4	.13765	.00055	.00061	4	.07176	.00016	.00018
5	.13768	.00062	.00079	5	.07177	.00018	.00023
10	.13780	.00078	.00102	10	.07181	.00022	.00029
20	.13793	.00084	.00100	20	.07184	.00024	.00028
30	.13801	.00084	.00096	30	.07186	.00024	.00027
40	.13806	.00083	.00092	40	.07188	.00024	.00026
50	.13810	.00082	.00089	50	.07189	.00023	.00025
100	.13820	.00077	.00080	100	.07192	.00022	.00023
200	.13828	.00071	.00072	200	.07194	.00020	.00020

(e); $n=160$			
c	(1)	(2)	(3)
1	.03666	.00002	-.00007
2	.03666	.00003	-.00002
3	.03666	.00004	.00003
4	.03667	.00004	.00005
5	.03667	.00005	.00006
10	.03668	.00006	.00008
20	.03669	.00006	.00008
30	.03670	.00006	.00007
40	.03670	.00006	.00007
50	.03670	.00006	.00007
100	.03671	.00006	.00006
200	.03672	.00005	.00005

4. Comparison of the Estimators under the Loss L_2 .

Haff [2] has shown that an estimator of Σ of order $p \times p$ of the form

$$\Sigma_{2H} = a \left(S + \frac{b}{\text{tr } S^{-1}} I \right), \quad a = 1/n, \quad 0 < b \leq \frac{2(p-1)}{n},$$

is uniformly better than the best multiple S/n of S under the loss function L_2 .

It has been proved by Sugiura and Fujimoto [ibid] that

$$(4.1) \quad R_2(\Sigma, \Sigma_{2H}) - R_2(\Sigma, aS) \\ = \frac{b}{n} \left[E \frac{\text{tr } \Sigma^{-1}}{\text{tr } S^{-1}} - n + \frac{bn}{2} E \frac{\text{tr } S^{-2}}{(\text{tr } S^{-1})^2} - \frac{b^2 n}{3} \frac{\text{tr } \Sigma^{-3}}{(\text{tr } \Sigma^{-1})^3} \right] + O(n^{-4}),$$

$$(4.2) \quad E \frac{\text{tr } \Sigma^{-1}}{\text{tr } S^{-1}} = n - p - 1 + 2E \frac{\text{tr } S^{-2}}{(\text{tr } S^{-1})^2}$$

and

$$(4.3) \quad E \frac{\text{tr } S^{-2}}{(\text{tr } S^{-1})^2} = \frac{\text{tr } \Sigma^{-2}}{(\text{tr } \Sigma^{-1})^2} + \frac{1}{n} \left[\frac{6(\text{tr } \Sigma^{-3})^2}{(\text{tr } \Sigma^{-1})^4} - \frac{8 \text{tr } \Sigma^{-3}}{(\text{tr } \Sigma^{-1})^3} + \frac{\text{tr } \Sigma^{-2}}{(\text{tr } \Sigma^{-1})^2} + 1 \right] + 0(n^{-2}).$$

An optimal value of b , in the sense that it minimizes an upper bound of $R_2(\Sigma, \Sigma_{2H}) - R_2(\Sigma, aS)$ is $b_0 = (p-1)/n$ and we shall take $b = b_0$ in the estimator Σ_{2H} .

Thus, for $p=2$, $\Sigma = \text{diag}(1, c)$ and $b = n^{-1}$, from (4.1) we get

$$(4.4) \quad R_2(\Sigma, \Sigma_{2H}) - R_2(\Sigma, aS) = \frac{3}{n^2} + \frac{5}{2n^2} \cdot \frac{1+c^2}{(1+c)^2} + \frac{5}{2n^3} \left[\frac{6(1+c^2)^2}{(1+c)^4} - \frac{8(1+c^3)}{(1+c)^3} + \frac{(1+c^2)}{(1+c)^2} + 1 \right] - \frac{(1+c^3)}{3n^3(1+c)^3} + 0(n^{-4}).$$

The Sugiura-Fujimoto estimator Σ_{2SF} dominating S/n is

$$(4.5) \quad \frac{1}{n} \left(S + b \frac{\text{tr } CS^{-1}}{\text{tr } (CS^{-1})^2} C \right), \quad 0 < b \leq 2(p-1)/n,$$

where C is an arbitrary positive definite matrix. We take $C=I$ and $b=(p-1)/n$. This value of b is optimal in the sense that it minimizes an upper bound of the risk difference

$$(4.6) \quad R_2(\Sigma, \Sigma_{2SF}) - R_2(\Sigma, S/n) = \frac{b}{n} \left[\left(\frac{nb}{2} - p - 1 \right) E \frac{(\text{tr } S^{-1})^2}{\text{tr } S^{-2}} - 2 + 4E \frac{\text{tr } S^{-3} \text{tr } S^{-1}}{(\text{tr } S^{-2})^2} \right] - \frac{b^3}{3} \cdot \frac{(\text{tr } \Sigma^{-1})^3 \text{tr } \Sigma^{-3}}{(\text{tr } \Sigma^{-2})^3}.$$

For (4.6), we have taken $C=I$ and b arbitrary, when $p=2$, $\Sigma = \text{diag}(1, c)$ and $b = n^{-1}$, it can be seen from (3.6) and (3.7) that

$$(4.7) \quad R_2(\Sigma, \Sigma_{2SF}) - R_2(\Sigma, S/n) = \frac{5(1+c)^2}{2n^2(1+c^2)} + \frac{4(1+c)(1+c^3)}{n^2(1+c^2)^2} - \frac{2}{n^2} - \frac{5}{2n^3} \left[8 \frac{(1+c^4)(1+c)^2}{(1+c^2)^3} - \frac{(1+c)^4}{(1+c^2)^2} - \frac{8(1+c^3)(1+c)}{(1+c^2)^2} - \frac{(1+c)^2}{(1+c^2)} + 2 \right] + \frac{4}{n^3} \left[\frac{24(1+c)(1+c^3)(1+c^4)}{(1+c^2)^4} - \frac{2}{(1+c^2)^3} \right] \times \{(1+c)^3(1+c^3) + 12(1+c)(1+c^5) + 4(1+c^5)^2\}$$

$$\begin{aligned}
 & + \frac{1}{(1+c^2)^2} \{(1+c)(1+c^3)+6(1+c^4)\} + 3 \frac{(1+c)^2}{1+c^2} \\
 & - \frac{(1+c)^3(1+c^3)}{3n^3(1+c^2)^3} + O(n^{-4}).
 \end{aligned}$$

The values of $R_2(\Sigma, \Psi_2)$, $R_2(\Sigma, \Sigma_{2H}) - R_2(\Sigma, \Psi_2)$ and $R_2(\Sigma, \Sigma_{2SF}) - R_2(\Sigma, \Psi_2)$ have been calculated from (2.13) and (4.7) and presented in Tables 2(a)-(e) for various values of $c \geq 1$ and n . We find that Ψ_2 is uniformly better than Σ_{2H} for all values of c and n under consideration. The estimator Ψ_2 is seen to be better than Σ_{2SF} for all c and $n=10$. However, it is worse than Σ_{2SF} at $c=1, 2$ and $n=20, 40, 80, 160$. For L_1 also one can expect the same result because $L_1/2 \rightarrow L_2$ as $n \rightarrow \infty$ (see, for example, Haff [2]). Incidentally, “ c greater than or equal to 1” is no restriction because of the scale equivariance of the estimators.

Tables 2(a)-(e). Asymptotic comparison of the estimators Ψ_2, Σ_{2H} and Σ_{2SF} . Columns (1), (2) and (3) give the values of $R_2(\Sigma, \Psi_2)$, $R_2(\Sigma, \Sigma_{2H}) - R_2(\Sigma, \Psi_2)$ and $R_2(\Sigma, \Sigma_{2SF}) - R_2(\Sigma, \Psi_2)$ respectively for $\Sigma = \text{diag}(1, c)$.

(a); $n=10$				(b); $n=20$			
c	(1)	(2)	(3)	c	(1)	(2)	(3)
1	.30525	.00356	.00597	1	.15084	.00077	-.00049
2	.30558	.00394	.00101	2	.15091	.00096	-.00012
3	.30605	.00448	.00201	3	.15103	.00119	.00078
4	.30649	.00496	.00408	4	.15113	.00137	.00141
5	.30688	.00533	.00558	5	.15122	.00151	.00179
10	.30829	.00621	.00792	10	.15155	.00181	.00232
20	.30982	.00641	.00783	20	.15191	.00191	.00228
30	.31070	.00623	.00732	30	.15211	.00188	.00216
40	.31129	.00601	.00688	40	.15225	.00184	.00206
50	.31173	.00581	.00653	50	.15235	.00179	.00198
100	.31296	.00507	.00546	100	.15263	.00163	.00173
200	.31397	.00432	.00452	200	.15286	.00147	.00152

(c); $n=40$				(d); $n=80$			
c	(1)	(2)	(3)	c	(1)	(2)	(3)
1	.07516	.00018	-.00037	1	.03753	.00004	-.00012
2	.07517	.00023	-.00008	2	.03754	.00006	-.00003
3	.07520	.00030	.00023	3	.03754	.00008	.00006
4	.07523	.00035	.00039	4	.03755	.00009	.00010
5	.07525	.00040	.00050	5	.03756	.00010	.00013
10	.07533	.00048	.00062	10	.03758	.00012	.00016
20	.07541	.00051	.00061	20	.03760	.00013	.00016
30	.07546	.00051	.00058	30	.03761	.00013	.00015
40	.07550	.00050	.00055	40	.03762	.00013	.00014
50	.07552	.00049	.00053	50	.03762	.00013	.00014
100	.07559	.00045	.00047	100	.03764	.00012	.00012
200	.07564	.00041	.00042	200	.03765	.00011	.00011

(e); $n=160$

c	(1)	(2)	(3)
1	.01876	.00001	-.00004
2	.01876	.00001	-.00001
3	.01876	.00002	.00002
4	.01876	.00002	.00003
5	.01876	.00003	.00003
10	.01877	.00003	.00004
20	.01877	.00003	.00004
30	.01878	.00003	.00004
40	.01878	.00003	.00004
50	.01878	.00003	.00004
100	.01878	.00003	.00003
200	.01879	.00003	.00003

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